

¹ Impact of hindcast length on estimates of seasonal ² climate predictability

Shi, W.¹, N. Schaller¹, D. MacLeod¹, T.N. Palmer^{1,2,3} and A. Weisheimer^{1,2,3}

¹Department of Physics, Atmospheric,
Oceanic and Planetary Physics, University
of Oxford, Oxford, OX1 3PU, UK

²Department of Physics, National Centre
for Atmospheric Science (NCAS),
University of Oxford, Oxford OX1 3PU, UK

³European Centre for Medium-Range
Weather Forecasts (ECMWF), Reading, UK

3 **Abstract**

4 It has recently been argued that single-model seasonal forecast ensembles are
5 overdispersive, implying that the real world is more predictable than indi-
6 cated by estimates of so-called perfect-model predictability, particularly over
7 the North Atlantic. However, such estimates are based on relatively short
8 forecast datasets comprising just 20 years of seasonal predictions. Here we
9 study longer 40-year seasonal forecast datasets from multi-model seasonal
10 forecast ensemble projects and show that sampling uncertainty due to the
11 length of the hindcast periods is large. The skill of forecasting the North At-
12 lantic Oscillation during winter varies within the 40-year datasets with high
13 levels of skill found for some sub-periods. It is demonstrated that whilst 20-
14 year estimates of seasonal reliability can show evidence of overdispersive be-
15 haviour, the 40-year estimates are more stable and show no evidence of overdis-
16 persion. Instead, the predominant feature on these longer timescales is un-
17 derdispersion, particularly in the tropics.

1. Introduction

18 There is no question that skilful seasonal forecasts can be made in the tropics (e.g.
19 *Barnston et al.* [2012]). However, the extent to which seasonal forecasts have useful in-
20 formation in the extratropics is more controversial. For example, whilst on the one hand
21 *Scaife et al.* [2014] recently showed that the new UK Met Office seasonal forecast model
22 GloSea5 was able to skilfully predict the wintertime North Atlantic Oscillation (NAO)
23 index for the period 1993-2012, on the other hand, *Weisheimer and Palmer* [2014] demon-
24 strated that seasonal predictions of temperature and precipitation were not reliable for
25 several regions in the extratropics, in particular over Europe.

26
27 Using the dataset of skilful NAO forecasts in GloSea5, *Eade et al.* [2014] suggested
28 that seasonal forecast ensembles created with initial condition uncertainty were under-
29 confident, or overdispersive, with too much noise in each ensemble. These would lead to
30 an unreasonably pessimistic estimate of seasonal predictability because the potential skill
31 would be underestimated implying that the real world would be more predictable than
32 the model world.

33
34 One of the difficulties with seasonal prediction research is that the sample size from
35 which estimates of forecast skill can be obtained is necessarily small: for start dates at a
36 given time of year, the sample size of seasonal forecasts for boreal winter from the 20-year
37 period 1992–2011, as used in *Eade et al.* [2014] is just 20. For example, one would hardly
38 implement changes to a numerical weather forecast model based on a sample of just 20

39 forecasts. Indications that 20 may be too small a number for robust estimates of skill can
40 be found in studies, e.g. by *Müller et al.* [2005] who showed that robust results for the
41 seasonal forecast skill of the NAO index were not stable with a sample size of 20, and by
42 *Kumar* [2009] demonstrating the effect on skill measures of small verification time series
43 due to sampling error.

44
45 In this paper we assess the *Eade et al.* [2014] claim that model estimates of extratropical
46 predictability in the North Atlantic region is unduly pessimistic. We analyse a consistent
47 set of seasonal forecast ensembles from a total of 8 individual models over various sub-
48 set of the 42-year period 1961–2001. It is found that several of these models have NAO
49 hindcast skill levels comparable to GloSea5. We then compute the ‘Ratio of Predictable
50 Components’ (*RPC*) diagnostic of *Eade et al.* [2014] to show that whilst individual model
51 ensembles can appear overdispersive over 20-year periods, they are not overdispersive over
52 40-year periods.

53
54 We conclude that the claims made in *Eade et al.* [2014] are consistent with sampling un-
55 certainty due to the limited length of the hindcast period. However, on 40-year timescales,
56 evidence suggests that single model ensembles are profoundly underdispersive. This
57 suggests that it remains crucially important to develop reliable methods to represent
58 parametrisation uncertainty (*Palmer* [2012], *Weisheimer et al.* [2014]).

2. Methodology

60 We use seasonal hindcast simulations over a 42-year period performed with 8 individual
61 model ensembles as part of in the European Union projects *DEMETER* (*Palmer et al.*
62 [2004]) and *ENSEMBLES* (*Weisheimer et al.* [2009]) to revisit the findings of *Eade et al.*
63 [2014] and to asses the predictability of the NAO. As is well known (*Hurrell et al.* [2001]),
64 the NAO is a mode of atmospheric variability over the North Atlantic region with wide-
65 ranging impacts on the weather and climate over Europe. In this paper, a simple index
66 of the NAO is defined following *Pavan and Doblas-Reyes* [2000] and *Doblas-Reyes et al.*
67 [2003], taking the model projections of the forecast anomalies of geopotential height at
68 500hPa (Z500) on the leading climatological Empirical Orthogonal Function. In addition,
69 we also computed an NAO index based on the normalised mean sea level pressure (MSLP)
70 difference between the Azores and Iceland.

71
72 The following models were used in our analysis: *D_ECMF* (ECMWF), *D_UKMO*
73 (MetOffice), *D_MEFR* (MétéoFrance) from the *DEMETER* system and *E_ECMF*
74 (ECMWF), *E_UKMO* (MetOffice), *E_KIEL* (IfM Kiel), *E_INGV* (INGV Bologna),
75 *E_MEFR* (MétéoFrance) from the *ENSEMBLES* system. The individual model ensem-
76 bles consist of 9 members that were created through perturbed initial conditions. For the
77 analysis we consider seasonal mean forecast anomalies for December to February (DJF)
78 from forecasts started on 1st November each year. The verification data were obtained
79 from the ERA40 Reanalysis Project (*Uppala et al.* [2005]).

80

Following *Eade et al.* [2014], ensemble-based estimates of predictability can be obtained from a diagnostic known as the ‘Ratio of Predictable Components’ (*RPC*) between the observed and model predicted values defined as

$$RPC = \frac{PC_{obs}}{PC_{mod}},$$

where PC is the predictable component in observations and in model hindcasts. In *Eade et al.* [2014], this is approximated by

$$RPC \geq \frac{r}{\sqrt{\sigma_{sig}^2 / \sigma_{tot}^2}}$$

81 where PC_{obs} is estimated directly from the explained variance given by the square of the
 82 correlation coefficient r between the ensemble mean model forecasts and the observations.
 83 The authors used the variance of the ensemble mean (σ_{sig}^2) relative to the average variance
 84 of individual ensemble members (σ_{tot}^2) to estimate PC_{mod} . For a perfect forecast system,
 85 the *RPC* should be close to 1. *RPC* values greater than one imply that the model is un-
 86 duly pessimistic in its estimate of skill, by being overdispersive. Conversely, *RPC* values
 87 below 1 point towards underdispersive and overconfident forecasts.

88
 89 However, such an interpretation has limitations. As discussed in *Kumar et al.* [2014],
 90 the definition of the model predictable component PC_{mod} depends on the particular fore-
 91 cast model used and cannot necessarily be indicative of the *true* potential predictability.
 92 Differences between actual skill levels (or PC_{obs} as estimated through the correlation r
 93 between the ensemble mean and observation) and potential skill of a perfect model (or
 94 PC_{mod} as estimated through the average correlation r_{perf} between the ensemble mean and
 95 the individual ensemble members) are related to errors in the model that lead to imperfect

96 biased forecasts. Furthermore, the above interpretation is only valid with a sufficiently
97 large hindcast length and ensemble size. With an insufficient sample size, estimates of
98 *RPC* can fluctuate above or below unity purely by chance, and no physical conclusions
99 can be reached about whether the ensemble system is under- or overdispersive overall.

100

101 Here we analyse the *RPC* of the NAO index and, similar to *Eade et al.* [2014], the global
102 MSLP fields simulated by the individual *DEMETER* and *ENSEMBLES* hindcasts. In or-
103 der to study the impact of the hindcast length on NAO skill and *RPC*, we analyse a large
104 number of combinations of hindcast years based on the full hindcast period 1960–2001.
105 Combinations of hindcast years were generated by randomly and independently sampling
106 from the very large number of all possible combinations of 5, 10, 15 . . . 40 years out of
107 the total 42-year period. For example, there exist 861 possible combinations of randomly
108 sampled 40 years. For shorter sub-periods there exist more conceivable combinations with
109 a maximum of more than 500 billion possible combinations for 20-year periods. In order
110 to have a comparable sample size for all considered sub-periods, our results are based on
111 20,000 draws from the combinations, with repetition.

112

113 We have tested the sensitivity and robustness of our results for longer hindcast periods
114 using two approaches: The first approach involved modifications of the random draws
115 of hindcast years by allowing resampling of years in each draw (with replacement). The
116 second approach is based on the finding that the maxima of the *RPC* distributions for
117 shorter hindcast periods up to 20 years can be approximated very well by an exponential

118 decay function, dependent on hindcast length. These exponential fits in turn provide an
119 alternative tool to extrapolate the RPC maxima for hindcast periods longer than 20 years.
120 While both approaches were found to result in some minor differences as to the exact shape
121 of the RPC distributions for long hindcast periods (not shown), the uncertainty ranges
122 from our resampling methodology as outlined above are consistent with these estimates.

123

3. Results

124 The NAO correlation coefficients between the ensemble mean and the verification data
125 for three different hindcast periods are given in Table 1 for the *DEMETER* and *ENSEM-*
126 *BLES* individual models for both the Z500-based and MSLP-based definitions of the NAO
127 index. Consistent with the results of *Müller et al.* [2005] it shows that there are differences
128 in the level of predictive skill between the two shorter sub-periods. This itself is indicative
129 that a 20-year period may be insufficient for a robust estimation of overall predictive skill.
130 The correlation between the modelled NAO indices and observations tends to be higher
131 for the late period 1980–2001 than for the early period 1960–1979. Some of the individual
132 models show significant correlations for the 20-year sub-periods (0.59 for *D_MEFR*, 0.45
133 for *D_ECMF* and 0.60 for *E_KIEL*). These levels of skill are comparable with the values
134 reported in *Scaiife et al.* [2014]. However, when we look at the entire 42-year hindcast
135 period 1960–2001, the correlations are considerably lower.

136

137 From the above described sampling algorithm, distributions of *RPC* values for the NAO
138 index (Z500 and MSLP) have been derived. Figure 1 shows these distributions for the

139 Z500-based index as box-and-whisker plots from the three *DEMETER* models *D_MEFR*,
140 *D_UKMO* and *D_ECMF* together with the two more recent versions of the ECMWF sea-
141 sonal forecast model: the version used in *ENSEMBLES* (identical to ECMWF's System
142 3) and the currently operational System 4 (for which only 30 years of hindcast data exist,
143 see also *Stockdale et al.* [2015]). For each ensemble the *RPC* distributions for different
144 lengths of hindcast data between 5 and 40 years (25 years in the case of System 4) are
145 displayed. As the length of the hindcast period increases, the *RPC* values for all models
146 decrease and the spread narrows. For the 5-year period, the *RPC* range includes both
147 large negative and large positive values. For the 20-year period, in particular, the upper
148 range of the *RPC* still clearly exceeds values of 1. Qualitatively very similar behaviour
149 was found for the analysis using either the MSLP-based NAO index or the other *ENSEM-*
150 *BLES* models, see figures in the Supplementary Material.

151

152 Since present-day operational seasonal forecast models are likely to be more skilful than
153 the typical *ENSEMBLES* models, as is the case of GloSea5, the upper range of the *RPC*
154 distribution will be more representative of possible *RPC* values from contemporary mod-
155 els. However, when 40 years of data are considered, no single model except for one gives
156 *RCP* values above 1 (the upper whisker of the distribution for *D_MEFR* just reaches 1).
157 Indeed, the entire distribution of *RPC* values for 40 years lies below 1 for all but one model.

158

159 Following *Eade et al.* [2014], Figure 2 shows global maps of *RPC* for mean sea level
160 pressure forecasts in DJF. The value of *RPC* shown is the maximum *RPC* for each grid-

161 point distribution based on 5,000 samples of possible combinations of hindcast years. For
162 clarity we only show the three individual *DEMETER* models for sub-periods of 5 (top
163 row), 20 (middle row) and 40 (bottom row) hindcast years. The corresponding figures
164 for the *ENSEMBLES* models can be found in the Supplementary Material. For a 5-year
165 sampling period, the maximum *RPC* is above 1 everywhere. When 20 years of hindcast
166 data are available, the maximum *RPC* in general decreases, with the tropics already indi-
167 cating values below 1. At 40 years, most of the regions of the world have maximum *RPC*
168 values that fall below 1. Contrary to *Eade et al.* [2014] this shows that when a sufficiently
169 long hindcast period is used so that the distribution of *RPC* converges, the seasonal fore-
170 cast ensembles from individual models are not underconfident (overdispersive) but rather
171 overconfident (underdispersive), in particular in the tropics.

172

4. Summary and Conclusion

173 In this study, we have investigated the seasonal forecast NAO skill during DJF in terms
174 of correlation between the ensemble mean and observations for a variety of seasonal fore-
175 cast models of the *DEMETER* and *ENSEMBLES* projects over different time periods.
176 For 20-year hindcast periods significant correlations for the NAO index were found, in
177 agreement with the results presented in *Scaife et al.* [2014]. However, no model produced
178 significant correlations throughout the entire 42-year period where seasonal hindcasts were
179 available (1960–2001).

180

181 In addition, we have analysed the ‘Ratio of Predictable Components’ (*RPC*) for sea-
182 sonal hindcasts of the NAO and mean sea level pressure in DJF. For periods of 20 years
183 or less, the distribution of possible *RPC* includes values greater than one which indicate
184 overdispersive conditions. However, for periods of 40 years, the maximum of all the dis-
185 tributions of *RPC* is always less than 1. This implies that the ensembles, if evaluated on
186 longer time scales, are not overdispersive. Indeed, by studying global surface pressure,
187 the overriding problem with single-model ensembles is their underdispersiveness in the
188 tropics. The interpretation of our results can lead to the conclusions that the findings
189 of *Eade et al.* [2014] merely suggest an inadequately small sample size where extreme
190 values of the *RPC* can easily be found above 1 when a 20-year sample size is used. These
191 extremes, however, all fall to values below 1 for the tested model hindcasts if a 40-year
192 hindcast period is used in the statistical analysis. Our results suggest that increasing the
193 sample size of GloSea5 by extending its current seasonal hindcasts length back to the
194 1960s would enable to test the robustness of GloSea5’s dispersion behaviour on longer
195 time scales.

196
197 Although our results suggest that single model ensembles are not overdispersive on av-
198 erage, it is still possible that current-generation climate models simulate a smaller range
199 of predictability than does the real world. This might occur if the real climate attractor
200 is more heterogeneous than the model attractor. That is to say, the real-world attractor
201 may have more distinct regions of stability and instability than does the more diffusive
202 climate-model attractor. Hence when the real world is evolving in a region of strong

203 predictability, the ensemble may be overdispersed and hence underconfident. Conversely,
204 when the real world is evolving in a region of weak predictability, the ensemble may be
205 underdispersed and hence overconfident. This is consistent with the notion that state-
206 space probability distributions for the real atmosphere show evidence of quasi-stationary
207 regimes, whilst simulated probability distributions, especially from low-resolution climate
208 models, tend to appear overly Gaussian (*Dawson et al.* [2012]).

209

210 The analysis of *Eade et al.* [2014] also studied ensemble forecasts from decadal pre-
211 diction experiments over the 46-year period 1960–2005. They concluded that the *RPC*
212 for mean sea level pressure over the globe is also underconfident (overdispersive), similar
213 to the seasonal forecast ensembles. However, the analysis is based on 4-year averages
214 of sea level pressure where the forecasts starting every year overlap in time. This im-
215 plies a large degree of dependence between the individual forecasts as there is considerate
216 overlap between the target forecast periods from different start years. Thus the effective
217 independent sample size of the hindcasts is not 46 but rather of the order of 10. By anal-
218 ogy, a similar situation would arise if one wanted to forecast the winter anomalies from
219 seasonal forecasts pooled together from several months of start dates across the autumn
220 and early winter of a given year; these forecasts cannot be counted as independent samples.

221

222 The number of members in the forecast ensemble is another source of sampling un-
223 certainty when estimating the correlation skill and *RPC*. While this study is focused on
224 the effect of the hindcast length, work to analyse how the ensemble size influences these

225 estimates is under way.

226

227 **Acknowledgments.** WS was funded through AOPP. This study was supported by
228 the EU projects SPECS (DM, TNP and AW) and EUCLEIA (NS and AW) funded by
229 the European Commission Seventh Framework Research Programme under the grant
230 agreement nos 308378 and 607085. TNP was supported by the European Research Council
231 Advanced Investigator Award ‘Towards the Prototype Probabilistic Earth-System Model’.

References

- 232 Barnston, A., M. Tippett, M. L’Heureux, S. Li, and D. DeWitt (2012), Skill of real-time
233 seasonal ENSO model predictions during 2002-11: Is our capability increasing?, *Bull.*
234 *Amer. Meteorol. Soc.*, *93*, 631–651, doi:DOI:10.1175/BAMS-D-11-00111.1.
- 235 Dawson, A., T. Palmer, and S. Corti (2012), Simulating regime structures in weather and
236 climate prediction models, *Geophys. Res. Lett.*, *39*, L21,805.
- 237 Doblas-Reyes, F. J., V. Pavan, and D. Stephenson (2003), The skill of multi-model sea-
238 sonal forecasts of the wintertime North Atlantic Oscillation, *Climate Dynamics*, *21*,
239 501–514, doi:10.1007/s00382-003-0350-4.
- 240 Eade, R., D. Smith, A. Scaife, E. Wallace, N. Dunstone, L. Hermanson, and N. Robinson
241 (2014), Do seasonal-to-decadal climate predictions underestimate the predictability of
242 the real world?, *Geophys. Res. Lett.*, *41*, 5620–5628, doi:10.1002/2014GL061146.
- 243 Hurrell, W., J., Y. Kushnir, and M. Visbeck (2001), The North Atlantic Oscillation,
244 *Science*, *291*, 601–603.

- 245 Kumar, A. (2009), Finite samples and uncertainty estimates for skill measures for seasonal
246 prediction, *Mon. Wea. Rev.*, *137*, 2622–2631.
- 247 Kumar, A., P. Peng, and M. Chen (2014), Is there a relationship between potential and
248 actual skill?, *Mon. Wea. Rev.*, *142*, 2220–2227.
- 249 Müller, W., C. Appenzeller, and C. Schär (2005), Probabilistic seasonal prediction of the
250 winter North Atlantic Oscillation and its impact on near surface temperature, *Climate
251 Dynamics*, *24*, 213–226, doi:10.1007/s00382-004-0492-z.
- 252 Palmer, T. (2012), Towards the probabilistic Earth-system simulator: a vision for the
253 future of climate and weather prediction, *Q.J.R. Meteorol. Soc.*, *138*, 841–861, doi:
254 DOI:10.1002/qj.1923.
- 255 Palmer, T., A. Alessandri, U. Andersen, P. Cantelaube, M. Davey, P. Delecluse, M. Deque,
256 E. Diez, F. Doblas-Reyes, H. Feddersen, R. Graham, S. Gualdi, J. Gueremy, R. Hage-
257 dorn, M. Hoshen, N. Keenlyside, M. Latif, A. Lazar, E. Maisonnavé, V. Marletto,
258 A. Morse, B. Orfila, P. Rogel, J. Terres, and M. Thomson (2004), Development of a
259 European multimodel ensemble system for seasonal-to-interannual prediction (DEME-
260 TER), *Bull. Amer. Meteorol. Soc.*, *85*(6), 853–872, doi:10.1175/BAMS-85-6-853.
- 261 Pavan, V., and F. Doblas-Reyes (2000), Multi-model seasonal forecasts over the Euro-
262 Atlantic: skill scores and dynamic features, *Climate Dynamics*, *16*, 611–625.
- 263 Scaife, A. A., A. Arribas, E. Blockley, A. Brookshaw, R. Clark, N. Dunstone, R. Eade,
264 D. Fereday, C. Folland, M. Gordon, L. Hermanson, J. Knight, D. Lea, C. MacLachlan,
265 A. Maidens, M. Martin, A. Peterson, D. Smith, M. Vellinga, E. Wallace, J. Waters, and
266 A. Williams (2014), Skillful long-range prediction of European and North American

- 267 winters, *Geophys. Res. Lett.*, *41*(7), 2514–2519, doi:10.1002/2014GL059637.
- 268 Stockdale, T., F. Molteni, and L. Ferranti (2015), Atmospheric initial conditions and the
269 predictability of the arctic oscillation, *Geophys. Res. Lett.*, *accepted*.
- 270 Uppala, S. M., P. Kallberg, A. Simmons, U. Andrae, V. Bechtold, M. Fiorino, J. Gibson,
271 J. Haseler, A. Hernandez, G. Kelly, X. Li, K. Onogi, S. Saarinen, N. Sokka, R. Allan,
272 E. Andersson, K. Arpe, M. Balmaseda, A. Beljaars, L. Berg, J. Bidlot, N. Bormann,
273 S. Caires, F. Chevallier, A. Dethof, M. Dragosavac, M. Fisher, M. Fuentes, S. Hage-
274 mann, E. Holm, B. Hoskins, L. Isaksen, P. Janssen, R. Jenne, A. McNally, J.-F. Mahfouf,
275 J.-J. Morcrette, N. Rayner, R. Saunders, P. Simon, A. Sterl, K. Trenberth, A. Untch,
276 D. Vasiljevic, P. Viterbo, and J. Woollen (2005), The ERA-40 re-analysis, *Q.J.R. Me-
277 teorol. Soc.*, *131*, 2961–3012, doi:doi: 10.1256/qj.04.176.
- 278 Weisheimer, A., and T. Palmer (2014), On the reliability of seasonal climate forecasts,
279 *J.R.Soc. Interface*, *11*(9620131162).
- 280 Weisheimer, A., F. Doblas-Reyes, T. Palmer, A. Alessandri, A. Arribas, M. Deque,
281 N. Keenlyside, M. MacVean, A. Navarra, and P. Rogel (2009), ENSEMBLES: A new
282 multi-model ensemble for seasonal-to-annual predictions: Skill and progress beyond
283 DEMETER in forecasting tropical Pacific SSTs, *Geophys. Res. Lett.*, *36*, L21,711, doi:
284 10.1029/2009GL040896.
- 285 Weisheimer, A., S. Corti, T. Palmer, and F. Vitart (2014), Addressing model error through
286 atmospheric stochastic physical parameterisations: Impact on the coupled ECMWF
287 seasonal forecasting system, *Phil. Trans. R. Soc. A*, *372*(201820130290).

Table 1. NAO correlations between model ensemble mean and observations based on Z500 (MSLP) for different hindcast periods. The first part shows results from the *ENSEMBLES* models. The second part shows results from the *DEMETER* models. Correlations where a *t*-test suggests significance at the 95% level are marked in bold.

	<i>E_ECMF</i>	<i>E_UKMO</i>	<i>E_KIEL</i>	<i>E_INGV</i>	<i>E_MEFR</i>
1960-1979	-0.16 (-0.35)	0.03 (0.17)	0.12 (0.60)	0.03 (-0.39)	0.07 (0.19)
1980-2001	0.20 (0.35)	0.02 (-0.08)	-0.07 (0.11)	0.22 (0.30)	0.35 (0.33)
1960-2001	0.07 (0.08)	-0.02 (0.00)	-0.08 (0.26)	0.10 (-0.02)	0.21 (0.26)
	<i>D_MEFR</i>	<i>D_ECMF</i>	<i>D_UKMO</i>		
1960-1979	0.26 (0.35)	-0.42 (0.10)	-0.05 (-0.47)		
1980-2001	0.59 (0.32)	0.45 (-0.05)	0.21 (0.02)		
1960-2001	0.38 (0.20)	-0.12 (-0.06)	-0.15 (-0.27)		

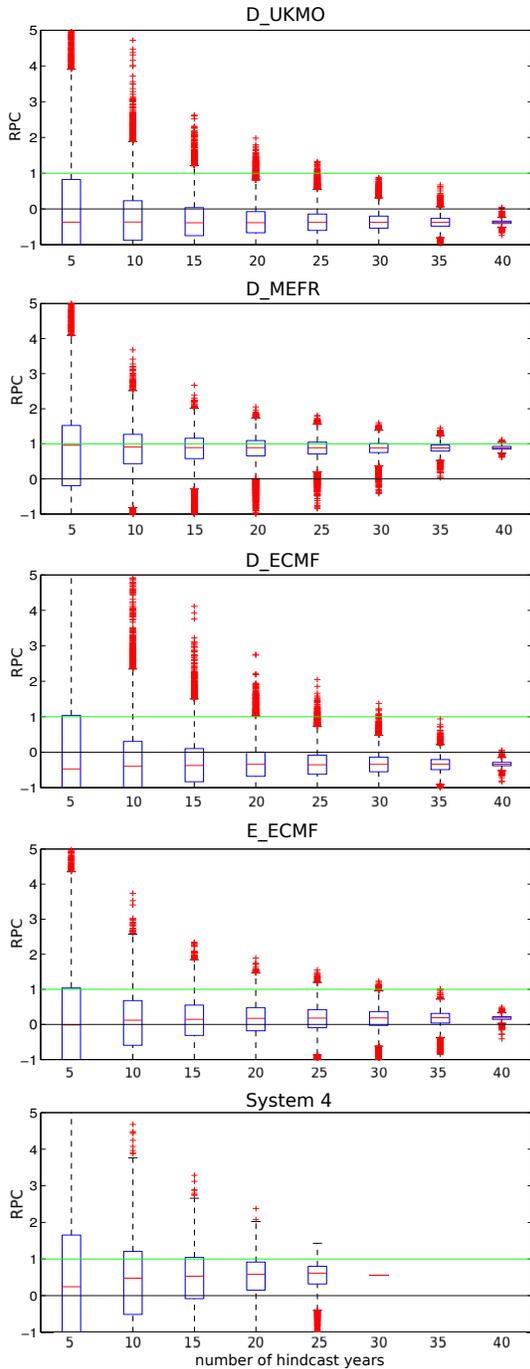


Figure 1. Box-and-whisker plots of the distributions of RPC of the $Z500$ -based NAO index in DJF as a function of the number of hindcast years used in the three *DEMETER* models (a–c), the *ENSEMBLES* model of ECMWF (d) and ECMWF’s currently operational forecasting System 4 (e). The box includes 50% of the data and the whiskers indicate approx. the 99% and 1% percentiles of the distribution. Outliers are marked with crosses.

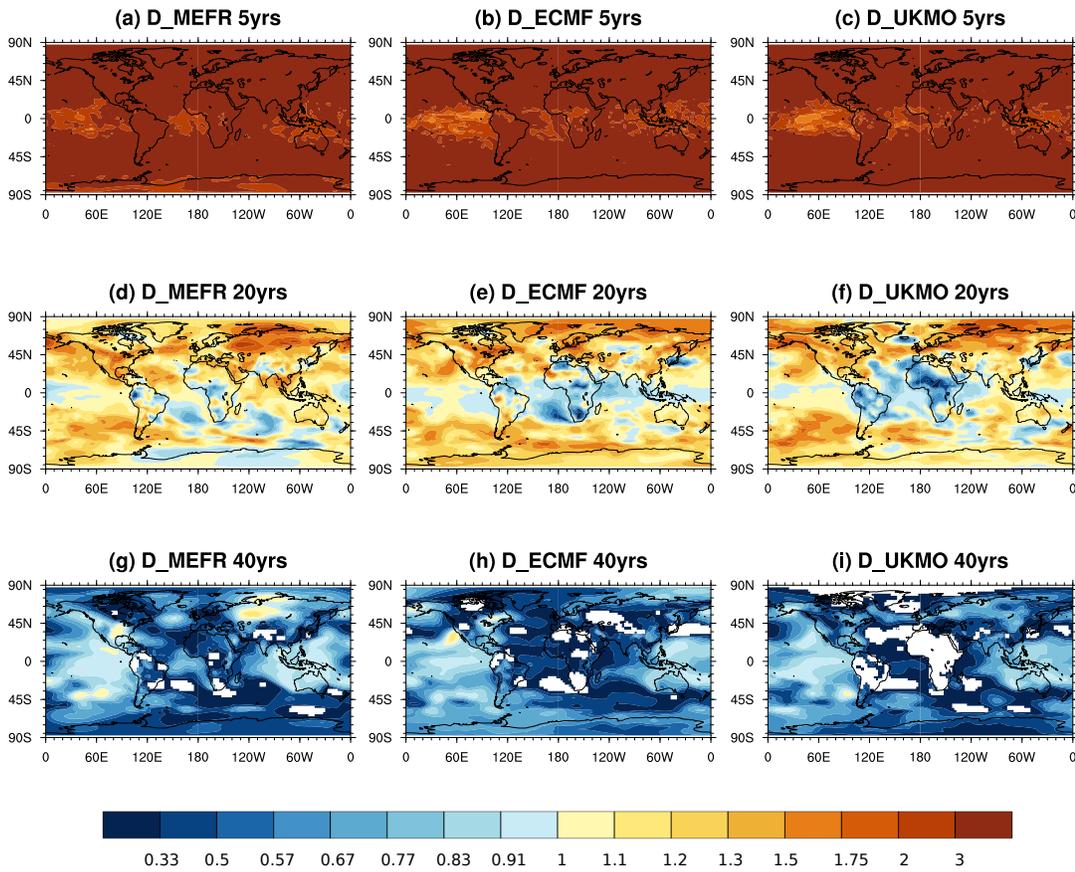


Figure 2. Maximum value of each *RPC* grid point distribution for mean sea level pressure in DJF. Results for 5-year, 20-year and 40-year time periods are shown for the three individual *DEMETER* model ensembles. Regions of negative *RPC* are masked out in white.